

## B.IX. Multivariate Extensions

Olivier Scaillet

University of Geneva and Swiss Finance Institute

# Outline

- 1 Introduction
- 2 Estimation of a VAR
- 3 Difficulties with Multivariate Extensions

# Introduction

Extension of AR process

⇒ Vector Autoregressive (VAR).

$Y_t$  = vector of dimension  $n$ .

$$Y_t = \mu + \Psi_1 Y_{t-1} + \dots + \Psi_p Y_{t-p} + \varepsilon_t.$$

Coefficients:  $\Psi_i = n \times n$  matrices.

Mean:  $\mu = n \times 1$  vector.

Error term:  $\varepsilon_t = n \times 1$  vector, usually taken as normally distributed.

# Introduction

Example:  $n = 2$ ,

$$\begin{pmatrix} Y_{1,t} \\ Y_{2,t} \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} + \begin{pmatrix} \psi_{1,11} & \psi_{1,12} \\ \psi_{1,21} & \psi_{1,22} \end{pmatrix} \begin{pmatrix} Y_{1,t-1} \\ Y_{2,t-1} \end{pmatrix} + \dots + \begin{pmatrix} \psi_{p,11} & \psi_{p,12} \\ \psi_{p,21} & \psi_{p,22} \end{pmatrix} \begin{pmatrix} Y_{1,t-p} \\ Y_{2,t-p} \end{pmatrix} + \begin{pmatrix} \varepsilon_{1,t} \\ \varepsilon_{2,t} \end{pmatrix}$$

$$\Rightarrow \begin{cases} Y_{1,t} = \mu_1 + \psi_{1,11} Y_{1,t-1} + \psi_{1,12} Y_{2,t-1} + \dots + \psi_{p,11} Y_{1,t-p} + \psi_{p,12} Y_{2,t-p} + \varepsilon_{1,t}, \\ Y_{2,t} = \mu_2 + \psi_{1,21} Y_{1,t-1} + \psi_{1,22} Y_{2,t-1} + \dots + \psi_{p,21} Y_{1,t-1} + \psi_{p,22} Y_{2,t-1} + \varepsilon_{2,t} \end{cases}$$

VAR has the form of a system of stacked linear regressions with the same set of explanatory variables in each regression.

# VAR Estimation

Maximum likelihood estimation under normality assumption will be numerically identical to estimation by OLS equation by equation.

⇒ Seemingly Unrelated REgressions (SURE) method.

Thus VAR are extremely easy to estimate since it only requires a regression software.

If we do not use the same set of explanatory variables, estimation by OLS equation by equation will be biased.

# Order of a VAR

The order of a VAR can be obtained by *likelihood ratio (LR) tests*.

Hypothesis of  $p_0$  lags against  $p_1$  lags with  $p_1 > p_0$ :

- 1 Estimate by OLS both models and derive the estimated residuals.
- 2 Compute the covariance matrices of residuals and the loglikelihoods

$$L(\hat{\Omega}_i) = (-Tn/2) \ln(2\pi) + (T/2) \ln(|\hat{\Omega}_i^{-1}|) - (Tn/2).$$

LR test statistic:

$$2 \left( L(\hat{\Omega}_1) - L(\hat{\Omega}_0) \right) \sim \chi^2(n^2(p_1 - p_0)).$$

# Goodness-of-Fit Criteria

We should balance good model fit and complexity (overfitting).

Choose the order  $p$  to minimise one of the following criteria:

- Akaike Information Criterion (AIC) - Tends to overestimate  $p$ :

$$C(p) = T \ln |\Omega| + 2(n^2 p).$$

- Bayesian Information Criterion (BIC):

$$C(p) = T \ln |\Omega| + (n^2 p) \ln(T).$$

- Hannan-Quin criterion(HQ):

$$C(p) = T \ln |\Omega| + 2(n^2 p) \ln(\ln(T)).$$

# Granger Causality

It helps to see whether a variable can help to forecast another one.

Bivariate example:

$$\begin{pmatrix} Y_{1,t} \\ Y_{2,t} \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} + \begin{pmatrix} \psi_{1,11} & \psi_{1,12} \\ \psi_{1,21} & \psi_{1,22} \end{pmatrix} \begin{pmatrix} Y_{1,t-1} \\ Y_{2,t-1} \end{pmatrix} + \begin{pmatrix} \varepsilon_{1,t} \\ \varepsilon_{1,t} \end{pmatrix}$$

We want to test whether  $Y_2$  does not cause  $Y_1$ , i.e.  $\psi_{1,12} = 0$ .

We need to regress both the general model

$$Y_{1,t} = \mu_1 + \psi_{1,11} Y_{1,t-1} + \psi_{1,12} Y_{2,t-1} + \varepsilon_{1,t}$$

and the restricted model

$$Y_{1,t} = \mu_1 + \psi_{1,11} Y_{1,t-1} + \eta_{1,t}$$

and compare the sum of squared residuals

$$\frac{T(SSR(\eta) - SSR(\varepsilon))}{SSR(\varepsilon)} \sim \chi^2(1).$$



# Difficulties

Lots of parameters to estimate, as illustrated by the following examples:

- VAR(1):

$$Y_t = \mu + \psi_1 Y_{t-1} + \varepsilon_t$$

where  $V[\varepsilon_t] = \Sigma$ .

⇒ Total number of parameters in this model is given by

$$n + n^2 + \frac{n(n+1)}{2}.$$

So, 5 assets = 45 parameters!

# Difficulties

- Multivariate GARCH(1,1):

$$\text{vech} (H_t) = A + B \text{vech} (H_{t-1}) + C \text{vech} (\varepsilon_{t-1} \varepsilon_{t-1}')$$

where

- vech stacks the lower-triangular part of a squared matrix (nxn) into a  $n(n+1)/2 + 1$  vector.
- $A$  = vector of size  $n(n+1)/2$ .  $B, C = (n(n+1)/2) \times (n(n+1)/2)$  matrices.

⇒ Total number of parameters:

$$\frac{n(n+1)}{2} + \frac{n^2(n+1)^2}{4} + \frac{n^2(n+1)^2}{4}.$$

5 assets = 465 parameters!!

# Difficulties

- Multivariate stochastic volatility:

$$Y_t = \mu + \exp(H_t/2)\varepsilon_t, V[\varepsilon_t] = \Sigma_\varepsilon$$

$$H_t = BH_{t-1} + \eta_t, V[\eta_t] = \Sigma_\eta$$

$\mu = n \times 1$  vector,  $B = n \times n$  matrix

⇒ Number of parameters:

$$n + n^2 + \frac{n(n+1)}{2} + \frac{n(n+1)}{2}$$

5 assets = 60 parameters!

# Possible Solutions

## Solutions to inflation of parameters:

- Impose some restrictions:
  - Diagonality on parameter matrices.
  - Symmetry on parameter matrices, etc